

Medical Statistics: Survival Analysis

Dr Nick Fieller

Department of Probability & Statistics

University of Sheffield

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Medical Statistics - Clinical Trials
Medical Statistics - Survival Analysis

3

Contents

Preliminaries

0: Introduction

1: Background & Basic Concepts

2: Single Sample Methods

2.1 Non-parametric Methods

2.6 Parametric Models

3: Two-Sample Comparisons

4: Regression Models

4.2 Parametric Regression Models

4.4 Proportional Hazards models



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Medical Statistics - Clinical Trials
Medical Statistics - Survival Analysis

2

Contents

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0: Introduction

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2.1 Non-parametric Methods

2.6 Parametric Models

3: Two-Sample Comparisons

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Medical Statistics - Survival Analysis

3

Organization of Course

◆ Two Components

- Clinical trials
 - Experiments on human (and animal) subjects
 - Ethical issues, efficient use of subjects, etc
- Survival Analysis
 - (analyzing data on length of lifetimes, e.g. times of remission in leukaemia)

◆ Approximately 10 lectures each



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4

Organization of course material

- ◆ Two sets of lecture notes (Survival & Clinical)
- ◆ Clinical Chapters 1 – 10 (~ 1 per lecture)
- ◆ Survival Chapters 1 – 4 main part of course
- ◆ Appendix 0: background maths
 - Maximum Likelihood Estimation
 - (but used only in a couple of places)
- ◆ Appendix 1 use of computer packages
 - SAS, SPSS, Minitab, S-PLUS
- ◆ Exercises & Task Sheets are in Course Booklet
 - Solutions follow later as appropriate



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5

Task Sheets & Exercises

◆ Task sheets:–

- ~ each week
- simple quick short exercises/reading
- reinforce / consolidate lecture material

◆ Exercises:–

- 3 sets during semester in weeks 5,8,10
- Work submitted within 2 weeks will be marked and returned

◆ See Study Guide

- recommendations on time to spend



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6



Task Sheets & Exercises

- ◆ Task sheets:–
 - are designed for you to test **your own understanding** of the course material
 - *you are responsible* for your own learning on the course — task sheets help you in self-assessment
- ◆ Exercises:–
 - **Prime route for individual feedback**
 - Task sheets often provide guide for exercises
- ◆ Unacceptable reasons for not submitting anything
 - I did not have enough time
 - I knew I could do them so I did not need to submit
 - I could not do anything so I did not think it was worth it

Solutions to Task Sheets & Exercises

- ◆ Exercises:–
 - Solutions available on web soon after submission
 - Printed solutions will be provided to those who submit
- ◆ Task sheets:–
 - are designed for you to test your own understanding of the course material
 - if necessary go back to lecture notes (etc) & re-read relevant sections
 - (and if necessary re-read again &)
 - Solutions will be provided on web pages in due course (for revision etc)
 - but **deliberately** these will not appear very quickly

Course web page

<http://nickfieller.staff.shef.ac.uk/>

- Click on  & then on [MAS6012/MAS461/MAS361 Medical Statistics](#)

- Lecture notes, task sheets, solutions & data sets available here after distribution in lectures
 - (I don't keep back copies)

Books

- ◆ **Campbell, M. J. (2001)**
Statistics at Square Two. BMJ
- ◆ **Collett, D. (2003)**
Modelling Survival Data in Medical Research. (2nd Edition) Chapman & Hall
- ◆ **Everitt, Brian & Rabe-Heskith, Sophia (2001)** *Analyzing Medical Data Using S-PLUS*. Springer. Support material at <http://web1.iop.kcl.ac.uk/loP/Departments/BioComp/splusBook.shtml>



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Preliminaries

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13

Objectives

- ◆ statistical modelling & analysis of *lifetime data*.
- ◆ Lifetime data arise especially in medical statistics and in reliability studies.
 - *survival time*:-
 - time from diagnosis to death of a patient
 - *time to recovery* or *remission* of a patient or
 - *time to failure* of an electronic component.



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14

Outline

- ◆ Types of survival data
- ◆ Censoring
- ◆ Parametric & non-parametric approaches

Single Sample Methods

- ◆ survivor & hazard functions
- ◆ Lifetables
- ◆ Kaplan-Meier estimators




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15

Two Sample Comparisons:-

- ◆ log rank test
- ◆ maximum likelihood test
- ◆ likelihood ratio test
- ◆ proportional hazards




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16

Regression Models

- ◆ parametric models
 - exponential & Weibull
- ◆ non-parametric methods
 - proportional hazards or Cox regression
 - partial likelihood




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17

Aims of survival analysis

- ◆ describe/model a single sample
 - making inferences on a single population
- ◆ compare two or more groups
 - effect of treatments on survival time
- ◆ investigate relationship with covariates
 - effects on survival time of covariates
 - adjust for covariates in comparisons



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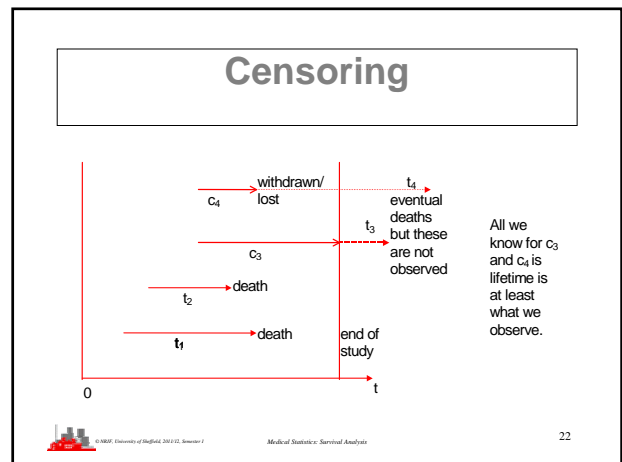
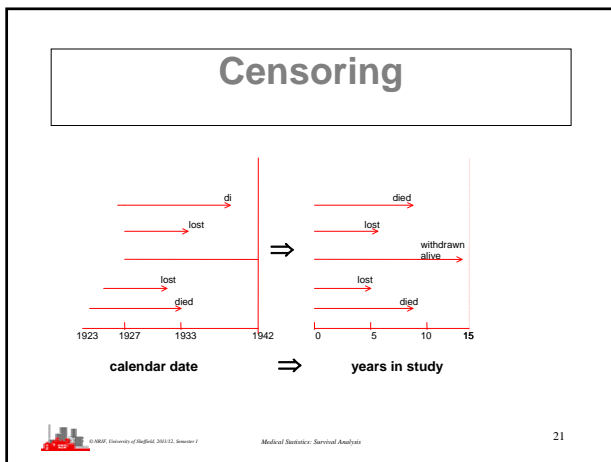
18



Contents

- Preliminaries
- 0: Introduction
- 1: **Background & Basic Concepts**
- 2: Single Sample Methods
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 - 2.6 Parametric Models
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- **Censoring**
 - ◆ event of interest not yet occurred
 - ◆ observe only survival time is **at least** t
 - ◆ some observations $t_i (i=1, \dots, n)$ are **censored**
 - ◆ censored observations provide information on survival so cannot be ignored




- ### Censoring
- **Right censoring**
 - ◆ lifetime exceeds some value
 - **Left censoring**
 - ◆ lifetime less than some value
 - **Interval censoring**
 - ◆ failure occurred during an interval

- **Aims:**
 - ◆ **estimate lifetime distributions**
 - \Rightarrow estimate properties of distribution
 - (median lifetime, prob of surviving > 5 years,...)
 - ◆ **Censoring**
 - non-parametric — lifetables, Kaplan-Meier
 - parametric — exponential, Weibull.




- survival time is a random variable T
- $T > 0$, continuous variable
 - ◆ p.d.f. $f(t)$ ($t > 0$)
 - ◆ d.f. $F(t) = P[T \leq t]$
 - ◆ (so $f(t) = F'(t)$ and $F(t) = \int_0^t f(u) du$)




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- **Survivor function**
 - ◆ $S(t) = P[T > t] = 1 - F(t)$
 - ◆ (so $S'(t) = -f(t)$, $S(t) = \int_t^\infty f(u) du$)
- **Hazard function**
 - ◆ $h(t) = \lim_{\delta t \rightarrow 0} \left[\frac{P[t \leq T < t + \delta t \mid T \geq t]}{\delta t} \right]$
 - ◆ "P[die at time T given survived until T]"



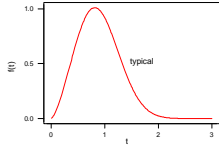
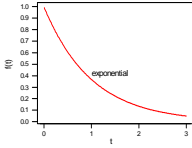
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- **Cumulative hazard function**
 - ◆ $H(t) = \int_0^t h(u) du = -\log_e S(t)$
 - ◆ $f(t)$, $S(t)$, $H(t)$ and $h(t)$ are equivalent characterizations of a survival distribution
 - ◆ all inter-related



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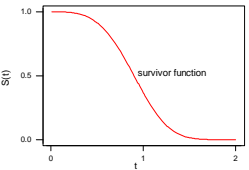

- **Typical patterns**



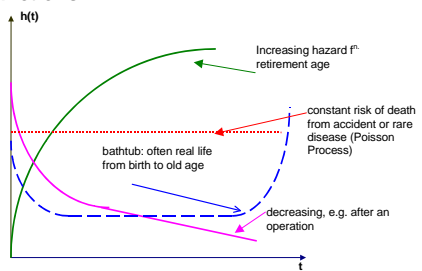
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- **Survivor function**

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Hazard functions:




Increasing hazard $h(t)$ retirement age

constant risk of death from accident or rare disease (Poisson Process)

bathtub: often real life from birth to old age


decreasing, e.g. after an operation



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- Choose appropriate family of models by recognizing form of hazard function
 - ◆ practical situation
 - ◆ initial investigation of the data
- then can estimate parameters in model



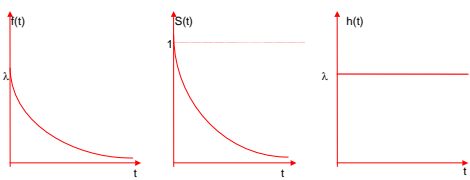

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- **Example:-** exponential
 - ◆ $f(t) = \lambda e^{-\lambda t}$
 - ◆ $S(t) = e^{-\lambda t}$
 - ◆ $h(t) = \lambda$ (**NB** constant)
 - $[= f(t)/S(t)]$
- **only** constant for exponential survival distribution




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- exponential

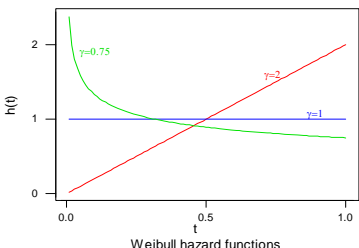

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- **Example:-**Weibull
 - ◆ $f(t) = \lambda \gamma t^{\gamma-1} \exp(-\lambda t^\gamma)$
 - ◆ $S(t) = \exp(-\lambda t^\gamma)$
 - ◆ $h(t) = \lambda \gamma t^{\gamma-1}$
 - $\gamma > 1$: increasing
 - $\gamma = 1$: constant (exponential)
 - $\gamma < 1$: decreasing



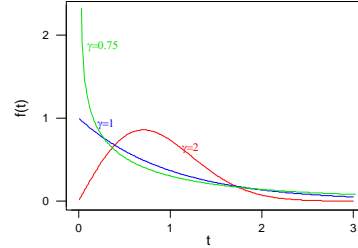

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- e.g. $\lambda = 1$; Weibull hazard functions

Weibull hazard functions

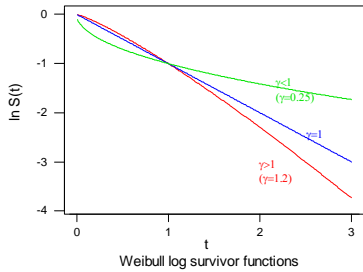
- e.g. $\lambda = 1$; Weibull density functions

Weibull Density Functions



- e.g. $\lambda = 1$; Weibull log-survivor functions



Weibull log survivor functions

- Weibull family
 - ♦ very flexible models
 - ♦ allows both increasing ($\gamma > 1$) & decreasing ($\gamma < 1$) hazard functions
- Difficult to estimate if γ close to 1
 - (numerical instability)



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43

- **Lifetables**
- ◆ 3 types :-
- **Population** (from census or survey)
 - **Cohort** (follow a group throughout lifetimes)
 - **Clinical** (or follow-up) survival pattern of specific group of individuals.
- 
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- **Example**
- ◆ every patient followed up after treatment either until death or up to the end of 1992
 - ◆ aim is to estimate probability of surviving for k years in separate steps:
 - first estimate probability of surviving another year given survived up to start of year, for each year
 - then multiply conditional probabilities together
- 
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- 45

Year of treatment	number treated	Number alive on each anniversary				
		1 st	2 nd	3 rd	4 th	5 th
1987	62	58	51	46	45	42
1988	39	36	33	31	28	
1989	47	45	41	38	73	
1990	58	53	48	115		
1991	42	40	173			
	248	232				




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46

▪ prob survive first year = $232/248 = 0.936$

Year of treatment	number treated	Number alive on each anniversary				
		1 st	2 nd	3 rd	4 th	5 th
1987	62	58	51	46	45	42
1988	39	36	33	31	28	
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1990	58	53	48	115		
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	248	232				




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47

▪ prob survive second year, given alive at start $173/(232 - 40) = 0.901$

Year of treatment	number treated	Number alive on each anniversary				
		1 st	2 nd	3 rd	4 th	5 th
1987	62	58	51	46	45	42
1988	39	36	33	31	28	
1989	47	45	41	38	73	
1990	58	53	48	115		
1991	42	40	173			
	248	232				



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48



- prob survive third year, given alive at start $115/(173 - 48) = 0.920$

Year of treatment	number treated	Number alive on each anniversary				
		1 st	2 nd	3 rd	4 th	5 th
1987	62	58	51	46	45	42
1988	39	36	33	31	28	
1989	47	45	41	38	73	
1990	58	53	48	115		
1991	42	40	173			
	248	232				

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Year after treatment	Prob. of surviving each year	Prob. of dying each year	Lifetable (per 1000)	
			Number alive on each anniversary	Number dying during each year
x	p_x	q_x	l_x	d_x
0	0.936	0.064	1000	64
1	0.901	0.099	936	93
2	0.920	0.080	843	67
3	0.948	0.052	776	40
4	0.933	0.067	736	49
5			687	

$0.901 \times 936 = 843$

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- n_x — number alive at start of $(x, x+1)$
- d_x — number dying in $(x, x+1)$
- estimate of conditional probability of dying in $(x, x+1)$, given alive at x , is $p_x = d_x/n_x$
- What if subjects are lost to follow-up?
 - (and not known if still alive or not)

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- w_x — number lost to follow-up
 - includes those who have disappeared (i.e. last report last year)
 - + 'withdrawn alive'
- assume withdrawals are uniformly spread over $(x, x+1)$
- adjusted number at risk $n'_x = n_x - 1/2 w_x$

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- adjusted number at risk $n'_x = n_x - 1/2 w_x$
- adjusted estimate of conditional probability of dying in $(x, x+1)$, given alive at x , is $p_x = d_x/n'_x$
- Then $n_{x+1} = n_x - d_x - w_x$

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Interval since operation years	Last reported during this interval	Living at start of interval	Adjusted number at risk	Estimated probability of death	Estimated probability of survival	% of survivors after x years	Estimate of p.d.f.	Estimate of hazard function	
x to x+1	Died d_x	withdrawn w_x	n_x	n'_x	q_x	p_x	l_x	$\hat{p}_{x+1/2}$	$\hat{h}_{x+1/2}$
0-1	90	0	374	374.0	0.2406	0.7594	100	0.241	0.274
1-2	76	0	284	284.0	0.2676	0.7324	75.9	0.203	0.309
2-3	51	0	208	208.0	0.2452	0.7548	55.6	0.136	0.279
3-4	25	12	157	151.0	0.1656	0.8344	42.0	0.070	0.181
4-5	20	5	120	117.5	0.1702	0.8298	35.0	0.059	0.186
5-6	7	9	95	90.5	0.0773	0.9227	29.1	0.023	0.080
6-7	4	9	79	74.5	0.0537	0.9463	26.8	0.014	0.055
7-8	1	3	66	64.5	0.0155	0.9845	25.4	0.004	0.016
8-9	3	5	62	59.5	0.0504	0.9496	25.0	0.013	0.052
9-10	2	5	54	51.5	0.0388	0.9612	23.7	0.009	0.040
10-	21	26	47	—	—	—	22.8	—	—

$120 = 157 - 25 - 12$

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Interval since operation years	Last reported during this interval	Living at start of interval	Adjusted number at risk	Estimated probability of death	Estimated probability of survival	% of survivors after x years	Estimate of p.d.f.	Estimate of hazard function
x to x+1	Died d_x	withdrawn w_x	n_x	n'_x	q_x	p_x	$\hat{p}_{x+\frac{1}{2}}$	$\hat{h}_{x+\frac{1}{2}}$
0-1	90	0	374	374.0	0.2406	0.7594	100	0.241
1-2	76	0	284	284.0	0.2676	0.7324	75.9	0.309
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6-7	4	9	79	74.5	0.0537	0.9463	26.8	0.055
7-8	1	3	66	64.5	0.0155	0.9845	25.4	0.016
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9-10	2	5	54	51.5	0.0388	0.9612	23.7	0.040
10-	21	26	47	—	—	—	22.8	—

$120 = 157 - 25 - 12$
 $117.5 = 120 - \frac{1}{2} \times 5$

- Estimated survivor function is $\hat{S}_x = p_0 p_1 \dots p_{x-1}$
- estimate of pdf is $\hat{f}_{x+\frac{1}{2}} := \hat{S}_x - \hat{S}_{x+1} = \hat{S}_x q_x$
- Estimate of hazard is $\hat{h}_{x+\frac{1}{2}} := \frac{2q_x}{1+p_x}$

- Estimate of hazard is $\hat{h}_{x+\frac{1}{2}} = \frac{\hat{f}_{x+\frac{1}{2}}}{\hat{S}_{x+\frac{1}{2}}} = \frac{\hat{S}_x q_x}{\frac{1}{2}(\hat{S}_x + \hat{S}_{x+1})} = \frac{2q_x}{1+p_x}$

noting $\hat{S}_{x+1} = p_x \hat{S}_x$

- assumed
 - withdrawals have same probability of death as non-withdrawals.
 - Is loss to follow-up connected with condition?
 - p_x and q_x constant over study
 - (estimated by combining data from several years)
- estimates are subject to sampling error:


$$\text{Var}(\hat{S}_x) = \hat{S}_x^2 \sum_{j=1}^{x-1} \frac{d_j}{n_j(n_j-d_j)}$$

- Clinical life tables suggest the form of the hazard function
- Lifetable methods take data in groups. Information is lost if actual lifetimes (perhaps censored) are available & have been grouped

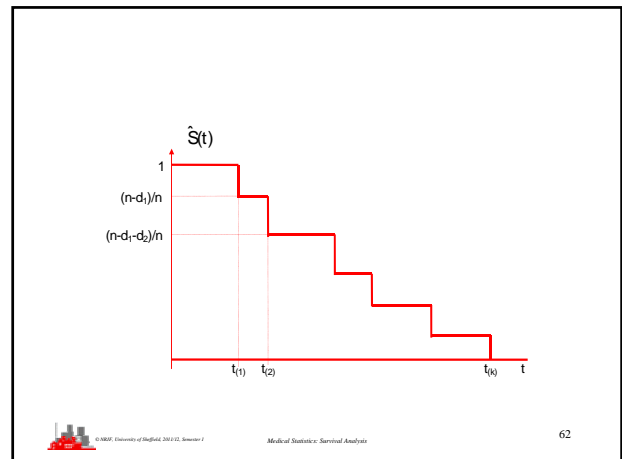
- Kaplan–Meier estimates
 - k ordered distinct lifetimes $t_{(1)} < t_{(2)} < \dots < t_{(k)}$
 - d_i — number of deaths at $t_{(i)}$ (so $\sum d_i = n$)



- $\hat{F}(t)$ = proportion of lifetimes < t
 $= \frac{1}{n} \sum_{j=1}^s d_j$ for $t_{(s)} \leq t < t_{(s+1)}$
- so $\hat{S}(t) = 1 - \hat{F}(t) = \frac{n - \sum_{j=1}^s d_j}{n}$
 for $t_{(s)} \leq t < t_{(s+1)}$




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- Let r_j be the number at risk (≡ number alive) just before $t_{(j)}$,
 Then $r_{j+1} = r_j - d_j$, so

$$\hat{S}(t) = \frac{n-d_1}{n} \cdot \frac{n-d_1-d_2}{n-d_1} \cdot \frac{n-d_1-d_2-d_3}{n-d_1-d_2} \dots \frac{n-d_1-d_2-\dots-d_s}{n-d_1-\dots-d_{s-1}}$$


$$= \left(1 - \frac{d_1}{r_1}\right) \left(1 - \frac{d_2}{r_2}\right) \dots \left(1 - \frac{d_s}{r_s}\right)$$



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$$= \prod_{j=1}^s \left(1 - \frac{d_j}{r_j}\right) \text{ for } t_{(s)} \leq t < t_{(s+1)}$$

- What if some observations censored?
 ♦ adjust numbers at risk r_j by allowing for censoring




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- $l_1, l_2, l_3, \dots, l_k$ numbers censored before time $t_{(j)}$



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$l_j = \#$ censored in the previous interval
 now $r_1 = n - l_1$; $r_{j+1} = r_j - d_j - l_{j+1}$ for $j=1, 2, \dots, k-1$
 [or $r_j = n - (d_1 + d_2 + \dots + d_{j-1}) - (l_1 + l_2 + \dots + l_j)$ for $j \geq 2$]
 \Rightarrow Kaplan-Meier product limit

$$\hat{S}(t) = \prod_{j=1}^s \left(1 - \frac{d_j}{r_j}\right) \text{ for } t_{(s)} \leq t < t_{(s+1)}$$


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- Notes
 - assumes that the l_j censored survive up to just after $t_{(j-1)}$ and then are removed
 - uncensored case is just a special case with $l_j=0$ all j
 - If $l_{k+1}>0$ then, since $r_k>d_k$

$$\hat{S}(t) = \prod_{j=1}^k \left(1 - \frac{d_j}{r_j}\right) > 0 \text{ when } t > t_{(k)}$$

But we know that $S(\infty) = 0$
 \Rightarrow Kaplan-Meier estimates are **biased** if maximum observation is censored

- $\hat{S}(t)$ is subject to sampling error

$$\text{var}(\hat{S}(t)) = [\hat{S}(t)]^2 \sum_{j=1}^s \frac{d_j}{r_j(r_j - d_j)} \text{ for } t_{(s)} \leq t < t_{(s+1)}$$
- so can get confidence bands for $S(t)$ with ± 2 x st.error

- estimate cumulative hazard $H(t)$ by

$$\hat{H}(t) = -\log_e(\hat{S}(t))$$
- or simpler is to use

$$\tilde{H}(t) = \sum_{j=1}^s \frac{d_j}{r_j} \text{ for } t_{(s)} \leq t < t_{(s+1)}$$

- Example
 - Remission times for 10 patients
 - 6 relapse 3.0, 6.5, 6.5, 10, 12, 15 months
 - 1 lost to follow-up at 8.4 months
 - 3 still in remission at end after 4.0, 5.7, 10.1 months
- i.e. 4 censored observations



j	t _(j)	l _j	r _j	d _j	Ŝ(t)		notes
					1	0 ≤ t < 3.0	
1	3.0	0	10	1	0.9	3.0 ≤ t < 6.5	9/10
2	6.5	2	7	2	0.643	6.5 ≤ t < 10.0	9/10 × 5/7
3	10.0	1	4	1	0.482	10.0 ≤ t < 12.0	9/10 × 5/7 × 3/4
4	12.0	1	2	1	0.241	12.0 ≤ t < 15.0	9/10 × 5/7 × 3/4 × 1/2
5	15.0	0	1	1	0	15 ≤ t	



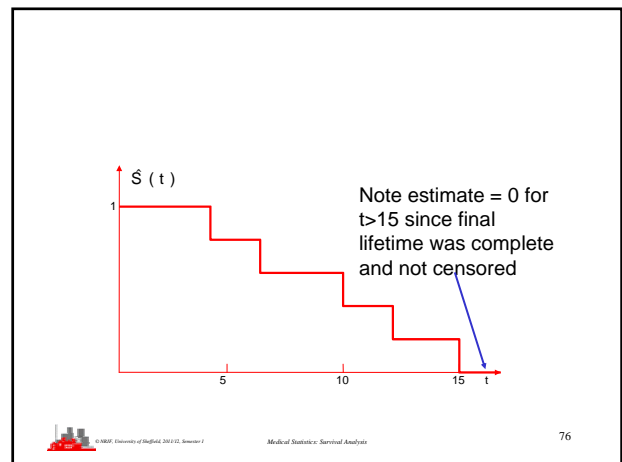
j	t _(j)	l _j	r _j	d _j	Ŝ(t)		notes
					1	0 ≤ t < 3.0	
1	3.0	0	10	1	0.9	3.0 ≤ t < 6.5	9/10
2	6.5	2	7	2	0.643	6.5 ≤ t < 10.0	9/10 × 5/7
3	10.0	1	4	1	0.482	10.0 ≤ t < 12.0	9/10 × 5/7 × 3/4
4	12.0	1	2	1	0.241	12.0 ≤ t < 15.0	9/10 × 5/7 × 3/4 × 1/2
5	15.0	0	1	1	0	15 ≤ t	

7 = 10 - 1 - 2



j	t _(j)	l _j	r _j	d _j	Ŝ(t)		notes
					1	0 ≤ t < 3.0	
1	3.0	0	10	1	0.9	3.0 ≤ t < 6.5	9/10
2	6.5	2	7	2	0.643	6.5 ≤ t < 10.0	9/10 × 5/7
3	10.0	1	4	1	0.482	10.0 ≤ t < 12.0	9/10 × 5/7 × 3/4
4	12.0	1	2	1	0.241	12.0 ≤ t < 15.0	9/10 × 5/7 × 3/4 × 1/2
5	15.0	0	1	1	0	15 ≤ t	

4 = 7 - 2 - 1



- Implementation in R:
 - ◆ Need to load library survival
 - > library(survival)
 - ◆ Need to create 'survival object' with Surv()
 - Surv(time, censor, type='right')
 - ◆ then use survfit() to estimate a survival curve
 - ◆ and plot() and summary() to see details




```

> library(survival)
Loading required package: splines
> load("tumour.Rdata")
> attach(tumour)
> tumour
      time censor
1    3.0      1
2    4.0      0
3    5.7      0
4    6.5      1
5    6.5      1
6    8.4      0
7   10.0      1
8   10.1      0
9   12.0      1
10  15.0      1
    
```

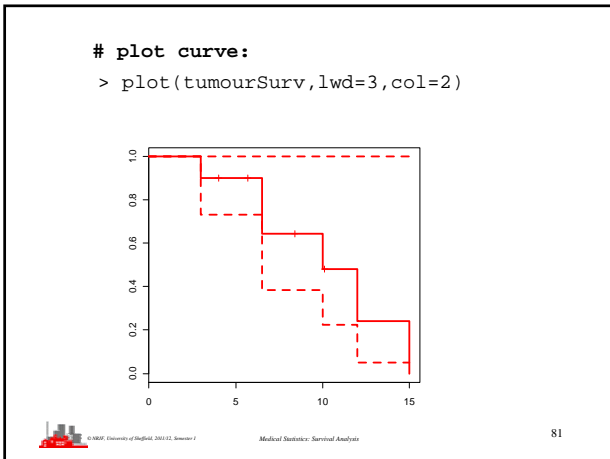
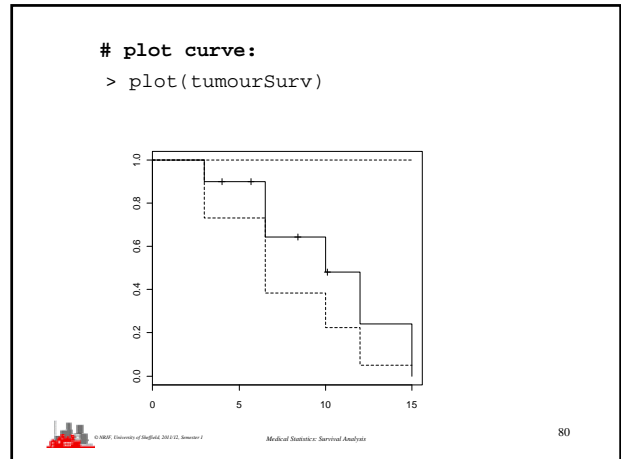



```
# create survival object in tumour.sv:
> tumour.sv <-Surv(time, censor, type = "right")
# estimate survival curve in tumourSurv:
> tumourSurv <-survfit(tumour.sv ~1, data=tumour)
# note 'regressing'/'relating' survival object on a
constant with ~ 1
# look at summary calculations
> summary(tumourSurv)
Call: survfit(formula = tumour.sv, data = tumour)
   time n.risk n.event survival std.err lower 95% CI upper 95% CI
3.0    10     1    0.900  0.0949  0.7320      1
6.5     7     2    0.643  0.1679  0.3852      1
10.0    4     1    0.482  0.1877  0.2248      1
12.0    2     1    0.241  0.1946  0.0496      1
15.0    1     1    0.000    NaN      NA      NA
```

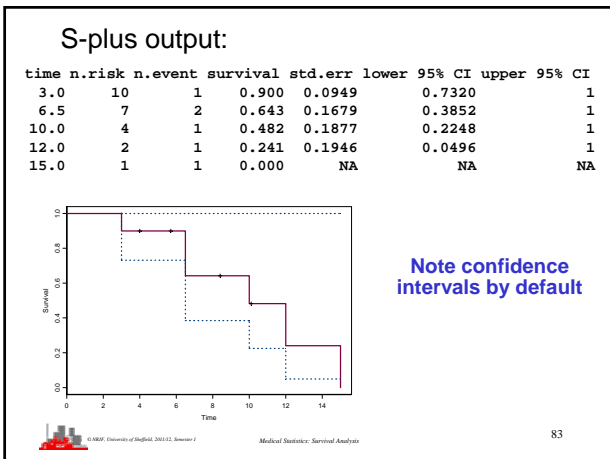
Note confidence interval for the K-M estimates and that in this small data set the interval is truncated at 1




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- S-plus implementation:
 - ◆ Statistics>Survival>Nonparametric Survival...
 - ◆ Need to create formula in dialogue box such as
Surv(time, censor, type='right')~1
- 
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- Minitab implementation:
 - ◆ Stat>Reliability/Survival>Nonparametric Dist Analysis-Right Censoring
 - ◆ Need to specify value indicating censored observations
- 
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Minitab output:

Kaplan-Meier Estimates

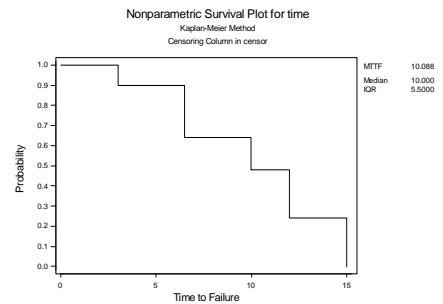
Time	Number at Risk	Number Failed	Survival Probability	Standard Error	95.0% Normal CI	
					Lower	Upper
3.0000	10	1	0.9000	0.0949	0.7141	1.0000
6.5000	7	2	0.6429	0.1679	0.3137	0.9720
10.0000	4	1	0.4821	0.1877	0.1142	0.8501
12.0000	2	1	0.2411	0.1946	0.0000	0.6225
15.0000	1	1	0.0000	0.0000	0.0000	0.0000



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Minitab graph:



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SPSS implementation:

- ◆ Analyze>Survival>Kaplan-Meier
- ◆ Need to specify value indicating **uncensored** values
- ◆ Note indication of censored values
 - Also indicated by S-PLUS



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SPSS Output:

Survival Analysis for TIME

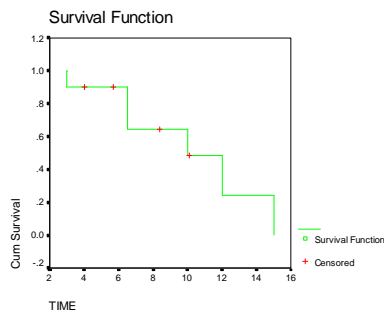
Time Remaining	Status	Cumulative Survival	Standard Error	Cumulative Events	Number
9	3	0	.9000	.0949	1
8	4	1			1
7	6	1			1
6	7	0			2
5	7	0	.6429	.1679	3
4	8	1			3
3	10	0	.4821	.1877	4
2	10	1			4
1	12	0	.2411	.1946	5
0	15	0	.0000	.0000	6



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SPSS Graph



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Summary

- ◆ Life tables
 - Grouped data
 - Allow for censoring by adjusting # at risk
- ◆ Kaplan-Meier
 - Individual data
 - Uncensored case = 1 – empirical CDF
 - Express as a product of terms $[1 - d_i/r_i]$
 - Censored case by adjusting # at risk



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Medical Statistics: Survival Analysis



- Kaplan-Meier estimates are the key exploratory tool for censored data
 - ◆ Cannot draw histograms of censored data
 - Don't know which bin to put a censored in
- Kaplan-Meier Plot is always step 1 in analysing censored data

