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### Discriminant Analysis

- **Techniques for when data classified into *known* groups**
  - ◆ Objectives:-
    - Effective data display utilising this extra information
    - Dimensionality reduction keeping information on group differences
    - assessment of group differences by examination of loadings (c.f. PCA)
    - Classification of new observations into the groups

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- **Methods:-**
  - ◆ **Linear Discriminant Analysis**
    - Based upon linear combinations of variables
    - Helps with all objectives
      - (cf PCA)
    - Can assess effectiveness of classification by simulation, jackknifing, randomization etc, (App 1)
    - **LDA is a *data analytic* method**
      - based just upon the data
    - (See Chapter 9 for Statistical Discrimination
      - Based on statistical models)
    - **LDA is a *supervised learning technique***

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- **If *primary objective* is *classification* of new observations other techniques may be better in **some** cases**
  - ◆ Quadratic Discriminant Analysis
    - See appendix 2
  - ◆ Neural Networks
    - See appendix 8
  - ◆ k-means classification
  - ◆ Nearest neighbour methods
  - ◆ .....&c.,&c.,.....

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- **Setup:-**
  - ◆ k **known** groups or categories  $G_i; i=1, \dots, k$
  - ◆  $n_i$  observations from each group,  $\sum n_i = n$
  - ◆ Each observation is p-dimensional
    - so data are  $\{x_{ij}; i=1, \dots, k; j=1, \dots, n_i\}$ ,  $x_{ij}$  a  $p \times 1$  vector)
    - data matrix for  $i^{\text{th}}$  group  $G_i$  is  $X_i'$  (so  $X_i'$  is  $n_i \times p$ )
    - $S_i$  is the within group  $G_i$  variance matrix ( $p \times p$ )
    - **W** is the **within groups variance**
    - **B** is the **between groups variance**
      - (see notes for formulæ)

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- **W** measures variability **inside** the groups
- **B** measures how **different** groups are
- Need to *compare* B with W
- Both W and B are  $p \times p$  matrices
  - ◆ i.e. variability measures are in p-dimensions
- Want to reduce the dimensionality
  - ◆ *Retaining* information on **B**  $\leftrightarrow$  **W** comparison

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- **Outline**
  - ◆ measure separation between known groups:-  $W^{-1}B$
  - ◆ 'compares' B in relation to W
  - ◆ In  $p=1$  dimension this is  $F = B / W$ 
    - The F-ratio in ANOVA
    - If F is big then more separation between groups (when compared with internal variability)
  - ◆ In  $p$  dimensions must assess how 'big'  $W^{-1}B$  is
    - This is a  $p \times p$  matrix — difficult, possibilities are determinant, largest eigenvalue, .... (see later)

- **Outline (ct<sup>d</sup>)**
  - Idea is project data into one dimension
    - ◆ (i.e. find linear combination of variables)
  - Which maximizes the F-statistic
    - ◆ (i.e. highlights distinction between groups)
  - Theorem shows this is achieved by first eigenvector of  $W^{-1}B$ 
    - i.e. corresponding to largest eigenvalue
      - strictly the **right** eigenvector ( $W^{-1}B$  not symmetric)

- **Outline (ct<sup>d</sup>)**
  - ◆ This is similar to the PCA theorem
    - (proof relies on **same procedure** as last time)
  - ◆ So, import analogies from PCA
    - Interpretation of coefficients of eigenvectors as loadings
    - Look also at 2<sup>nd</sup>, 3<sup>rd</sup>, ... eigenvectors of  $W^{-1}B$
    - Scree plots of eigenvalues
    - Plot data on eigenvectors (or **Crimcoords**)
      - even though eigenvectors are not orthogonal

▪ (end of outline)

### A Procedure for Maximization

- ◆ 1: Introduce some constraint
- ◆ 2: Introduce a Lagrange multiplier & define a new objective function
- ◆ 3: Differentiate w.r.t.  $x$  and set =0
- ◆ 4: Recognise this is an eigenequation with the Lagrange multiplier as eigenvalue
- ◆ 5: Deduce that there are **ONLY** a limited number of possible values for this eigenvalue (all of which can be calculated numerically)
- ◆ 6: Use some extra step to determine which eigenvalue gives the maximum (typically use the constraint somewhere)

- **Derivation of Crimcoords**
  - ◆ Project data by vector  $a_1$ 
    - i.e. take line combination of variables given by  $a_1$
  - ◆ so data sets for the  $k$  groups are  $X_i' a_1$
  - ◆ between groups variance is  $a_1' B a_1$
  - ◆ within groups variance is  $a_1' W a_1$
  - ◆  $W^{-1}B$  becomes
 
$$F_1 = \frac{a_1' B a_1}{a_1' W a_1}$$
  - ◆ So need to choose  $a_1$  to maximize  $F_1$

- Note that if we multiply  $a_1$  by a scalar constant  $c$  then  $F_1$  does not alter
 
$$F_1 = \frac{a_1' B a_1}{a_1' W a_1} = \frac{(ca_1)' B (ca_1)}{(ca_1)' W (ca_1)}$$

for any scalar  $c$

  - ◆ So can impose a **scale constraint** on  $a_1$ 
    - Could take  $a_1' a_1 = 1$
    - Simpler to take  $a_1' W a_1 = 1$ 
      - (makes differentiating  $F_1$  easier)
    - Maximize  $F_1$  subject to constraint  $a_1' W a_1 = 1$
    - i.e. maximize  $a_1' B a_1$  subject to constraint  $a_1' W a_1 = 1$



- Introduce a Lagrange multiplier  $\lambda_1$ :  
let  $\Omega_1 = a_1'Ba_1 - \lambda_1(a_1'Wa_1 - 1)$   
so  $\frac{\partial \Omega_1}{\partial a_1} = 2Ba_1 - 2\lambda_1Wa_1 = 0$
- i.e.  $W^{-1}Ba_1 - \lambda_1a_1 = 0$
- i.e.  $a_1$  is eigenvector of  $W^{-1}B$  corresponding to  $\lambda_1$ 
  - strictly a *right* eigenvector since  $W^{-1}B$  not symmetric

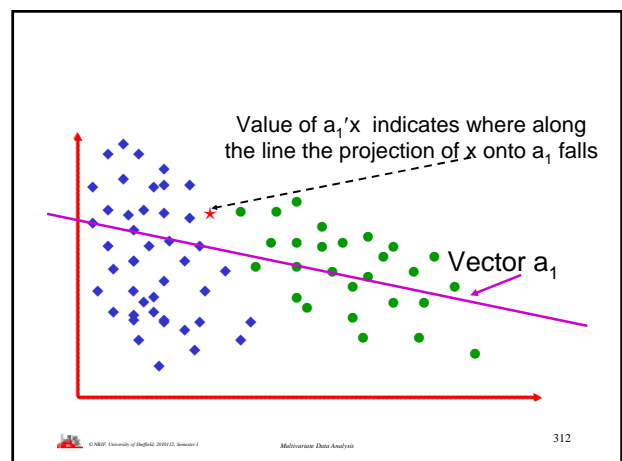
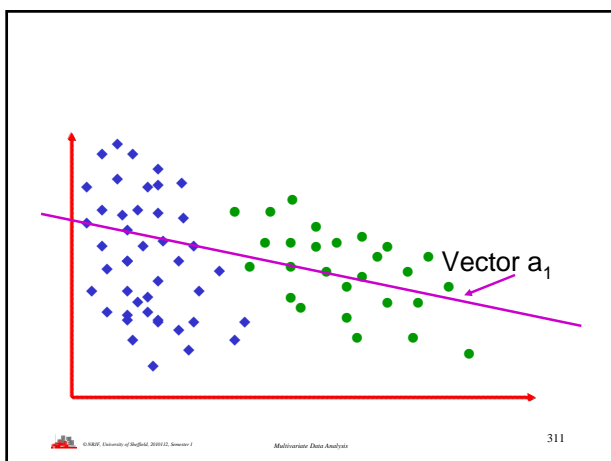
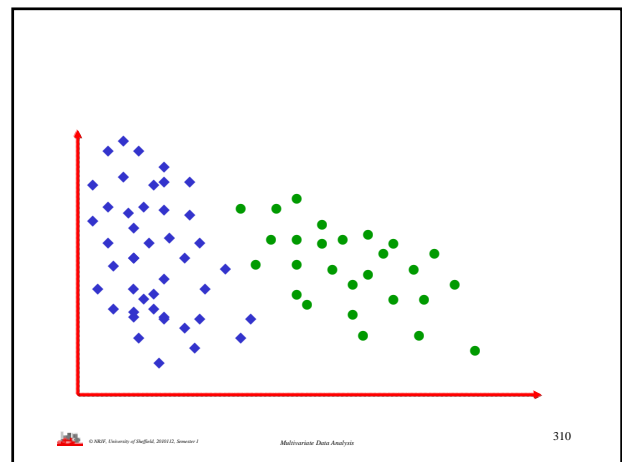
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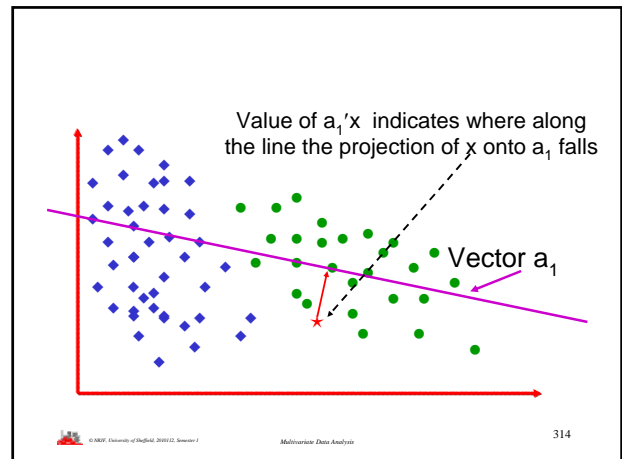
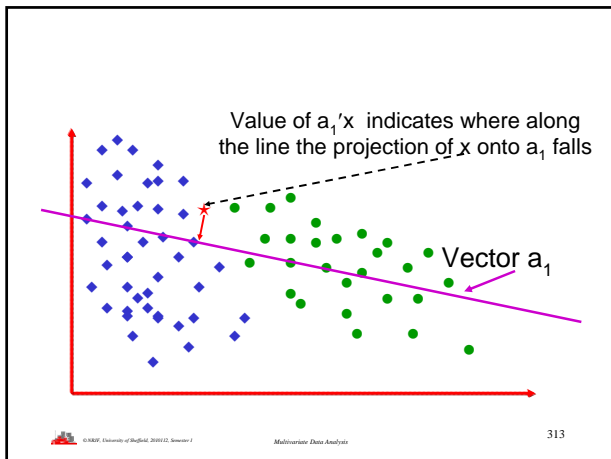
- Which eigenvalue is  $\lambda_1$ ?  
 $W^{-1}Ba_1 - \lambda_1a_1 = 0$ 
  - pre-multiply by  $a_1'W$
- so  $a_1'Ba_1 = \lambda_1a_1'Wa_1 = \lambda_1$ 
  - (since  $a_1'Wa_1 = 1$  — the original constraint)
  - **SO** to maximize  $a_1'Ba_1$  with  $a_1'Wa_1 = 1$  take  $\lambda_1$  as the *largest* eigenvalue of  $W^{-1}B$  &  $a_1$  as the corresponding eigenvector
    - (aside: compare Exercises 1, Q3)

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- **Fisher's Linear Discriminant Function**
  - ◆ The function of the p-vector  $x$   
 $f(x) = a_1'x$   
is known as *Fisher's linear discriminant function*
  - ◆ The value of this is the most sensitive to which group or category  $x$  belongs
  - ◆ to classify a new observation  $x$  a rule based on the value of  $f(x)$  is a leading contender
    - (but not the only contender)
    - (see later)

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■ ? Other eigenvalues of  $W^{-1}B$  ?  
 ◆  $W^{-1}B$  may have non-zero eigenvalues  $\lambda_1, \lambda_2, \dots, \lambda_r$ 

- where  $r = \text{rank}(W^{-1}B) \leq \min(k-1, p)$
- (= unless collinearities in group means)
- (or in data points inside each group)
  - or linear dependence between variables

 ◆ Only  $\lambda_1$  and  $a_1$  have strict interpretations  
 ◆ What can we do with the other eigenvalues and eigenvectors?

■ Note eigenvectors of  $W^{-1}B$  non-orthogonal
 

- $W^{-1}B$  is not symmetric (though both  $W$  and  $B$  are)
- ◆ However (for  $i^{\text{th}}$  &  $j^{\text{th}}$ ):
 
$$W^{-1}Ba_i - \lambda_i a_i = 0 \Rightarrow Ba_i - \lambda_i Wa_i = 0$$
 & 
$$W^{-1}Ba_j - \lambda_j a_j = 0 \Rightarrow Ba_j - \lambda_j Wa_j = 0$$
 so 
$$a_j' Ba_i - \lambda_i a_j' Wa_i = 0$$
 & 
$$a_i' Ba_j - \lambda_j a_i' Wa_j = 0$$
- ◆  $W$  &  $B$  symmetric;
  - so  $a_j' Ba_i = a_i' Ba_j$  &  $a_j' Wa_i = a_i' Wa_j$  and  $\lambda_i \neq \lambda_j$
 so 
$$a_j' Wa_i = 0 \text{ for } i \neq j$$

■  $a_i' Wa_j = 0$  for  $i \neq j$ 

- ◆  $a_i$  &  $a_j$  are 'orthogonal wrt  $W$ '
- ◆ The within groups variance  $W$  reflects the different variances and covariances of the variables
- ◆ If  $W = \sigma^2 I_p$ 
  - (i.e. all variables independent variance  $\sigma^2$ )
 then  $a_i' a_j = 0$  and  $a_i$  &  $a_j$  are orthogonal
- ◆ If  $W = D = \text{diag}(\sigma^2_{kk})$  [a diagonal matrix]
  - (i.e. all variables independent, different variances)
 then  $a_i' a_j = \sum \sigma^2_{kk} a_{ik} a_{jk} = 0$
- ◆  $W$  'corrects' for the variance structure of the variables

■ **Definitions:**

- ◆ The functions  $a_i'x$ ,  $i=1,2,\dots,r$  are called the **discriminant coordinates** or **crimcoords**
- ◆ The space spanned by [the first  $t$  of] them is called the **discriminant space**
- ◆ choose suitable  $t$  by **screen plot** of  $\lambda_1, \lambda_2, \dots$ 
  - (by analogy with PCA)
  - but don't have 'amount of discrimination' partitioned into 'separate bits'
    - contradistinction to PCA



- Traditional to plot crimcoords as if they were on orthogonal axes
- This 'distorts' the original space
  - ◆ but in a sensible way
  - ◆ equivalent to a form of 'average standardization' of the variables to bring them to a common scale of measurement
    - (i.e. common st.devs and 'orthogonalized')
    - Note:- common  $\equiv$  *within groups* (averaged over groups)

- ◆ transforming to discriminant space appropriate for displaying relationships & distinctions between groups
  - (allowing for the different scales of the variables and their structure)
- ◆ Given this, examining loadings of variables in crimcoords indicates which variables provide discrimination between groups
  - c.f. interpretation of principal components

### Computation

- ◆ S-plus:- function `lda( . )` in MASS library
  - full facility: `help(lda)` to find out more
- ◆ Minitab:
  - **Stat>Multivariate>Discriminant Analysis** **BUT cannot produce plots on crimcoords**
  - **Can** produce crimcoords 'by hand' from command line using spectral decomposition of  $W$  & eigenanalysis of  $B$ 
    - Cumbersome
  - Or use a *slick trick*:
    - cut&paste from **MANOVA** output

### Implementation in R

- ◆ Use command line function `lda( )` in MASS library
  - Modern Applied Statistics with **S**
    - » By Venables & Ripley
- ◆ syntax: `lda(group~V1+V2+...+Vp)`

- syntax: `lda(group~V1+V2+...+Vp)`
- Coordinates on crimcoords obtained in `predict.lda(lda(group~V1+V2+...+Vp))$x`
- i.e. data matrix multiplied by matrix of eigenvectors of  $W^{-1}B$ 
  - can then plot this to obtain plot on crimcoords
- To get new data multiplied by this matrix: `predict.lda(lda(group~V1+V2+...+Vp), newdata)$x`
- Gives new data matrix multiplied by this
  - best to store objects at each stage

### Example:- Anderson's Iris Data

- ◆ Historical note:-
  - Edgar Anderson consulted Ronald Fisher about the problem which motivated Fisher to develop the methodology of LDA ~70 yrs b.p.
  - most over-analyzed data set in world
  - if a technique doesn't work on the Iris Data then technique is 'useless'



```
> attach(irisnf)
> library(MASS)
> iris.lda<-
  lda(Variety~Sepal.l+Sepal.w+Petal.l+Petal.w)
> iris.crimcoord<-predict.lda(iris.lda)$x
> iris.lda
> plot(iris.lda)
> plot(iris.crimcoord)
```

```
> attach(irisnf)
• Attach data set
> library(MASS)
• Attach library containing lda(.) function
> iris.lda<-
+ lda(Variety~Sepal.l+Sepal.w+Petal.l+Petal.w)
• Perform linear discriminant analysis
and store in iris.lda
> iris.crimcoords<-predict.lda(iris.lda)$x
• Store coordinates in iris.crimcoords
> iris.lda
• Print everything in object iris.lda
```

```
> iris.lda ← print everything in object iris.lda
Call:
lda(Variety ~ Sepal.l + Sepal.w + Petal.l + Petal.w)

Prior probabilities of groups:
      1      2      3
0.3333333 0.3333333 0.3333333

Group means:
  Sepal.l Sepal.w Petal.l Petal.w
1  5.006  3.428  1.462  0.246
2  5.936  2.764  4.260  1.326
3  6.588  2.974  5.552  2.026

Coefficients of linear discriminants:
      LD1      LD2
Sepal.l -0.8129912  0.006189401
Sepal.w -1.5569539  2.157270079
Petal.l  2.1987600 -0.919346099
Petal.w  2.8266164  2.825877950

Proportion of trace:
      LD1      LD2
0.9912 0.0088
```

Eigenvectors of  $W^{-1}B$  corresponding to non-zero eigenvalues: elements give loadings of variables in each crimcoord.

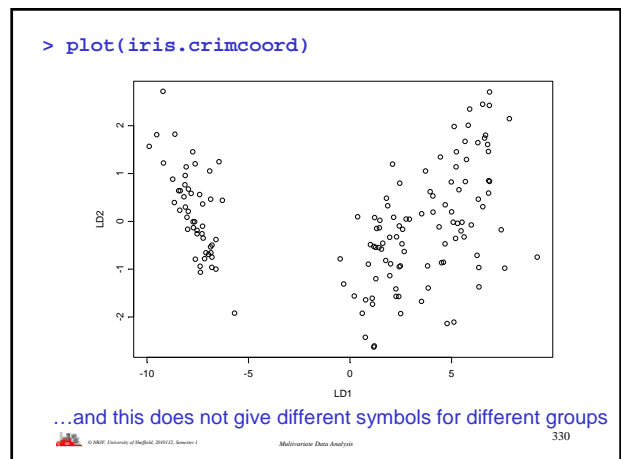
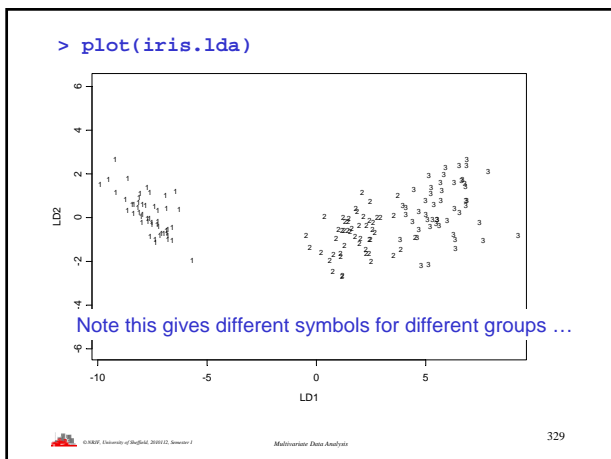
Eigenvalues as a percentage of total

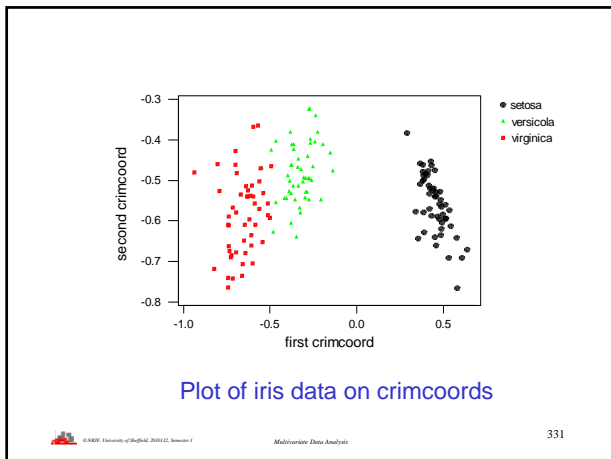
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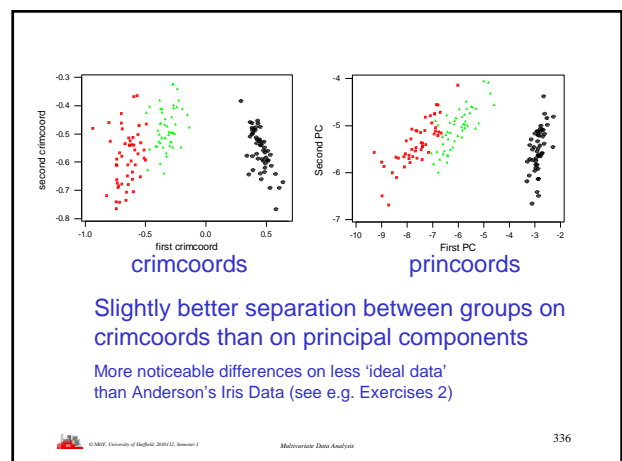
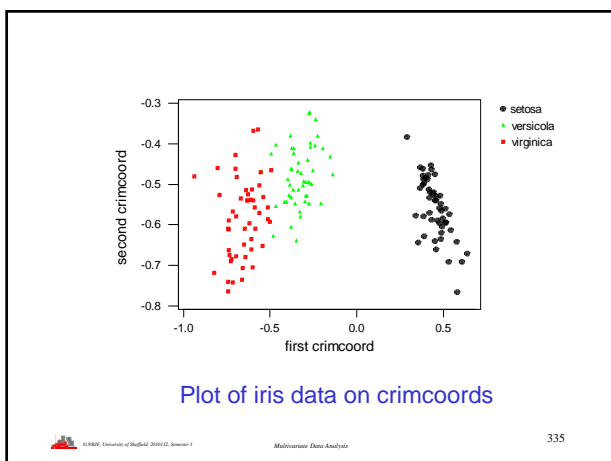
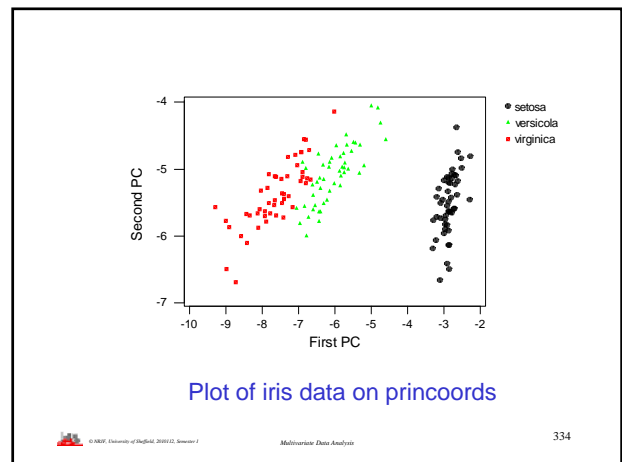
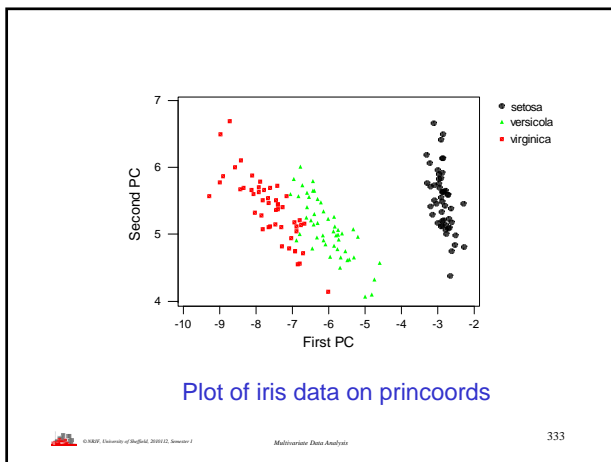
Eigenvalue 1      2
sepal-l  0.0671 -0.0005
sepal-w  0.1284 -0.1779
petal-l -0.1814  0.0758
petal-w -0.2331 -0.2331
```

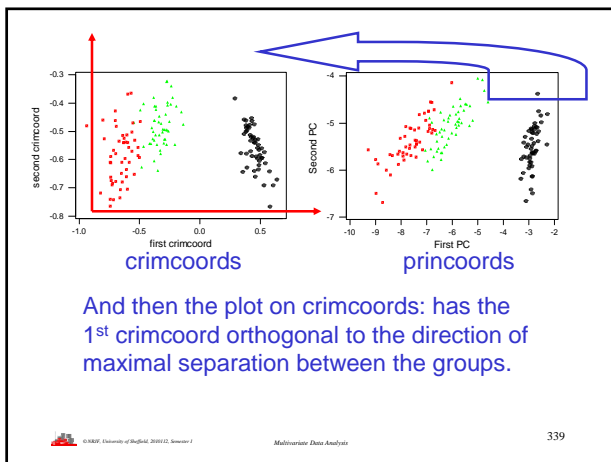
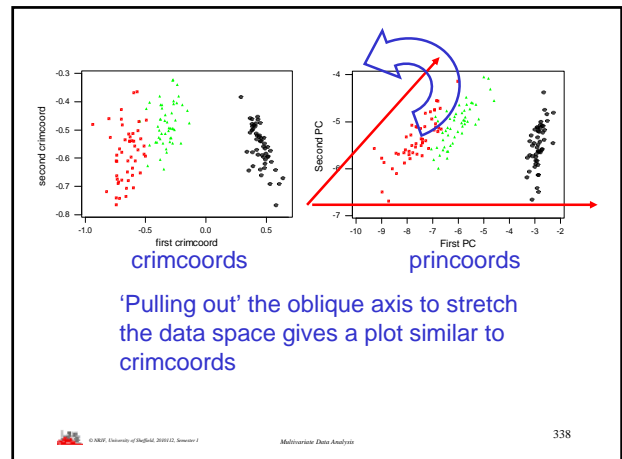
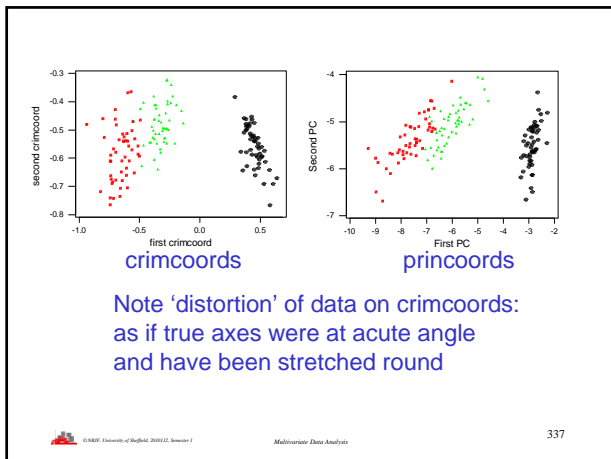
Numerically different from values produced by Minitab but only by a scale factor  $\Rightarrow$  same interpretation





- Slightly better separation than plots on principal components
  - Note that signs of principal components and crimcoords are arbitrary
  - Can flip plot to mirror image by
    - ◆ `MTB > let 'second PC' = - 'second PC'`
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- Minitab:
    - ◆ MANOVA: Multivariate Analysis of Variance
      - for statistical testing of differences between groups
        - (see next chapter)
      - option available for producing eigenvectors (& values) of  $W^{-1}B$
    - ◆ ANOVA>Balanced MANOVA...
      - followed by checking the 'Eigen analysis' box on the Results... menu
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◆ ANOVA>Balanced MANOVA...

- followed by checking the 'Eigen analysis' box on the Results... menu

◆ produces the commands for the eigenanalysis of the  $W^{-1}B$

```
MTB > ANOVA 'Sepal-1'-'Petal-w' = Variety;
SUBC> MANOVA;
SUBC> Eigen;
SUBC> NoUnivariate.
```

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◆ ANOVA>Balanced MANOVA...

- followed by checking the 'Eigen analysis' box on the Results... menu

◆ produces the commands for the eigenanalysis of the  $W^{-1}B$

```
MTB > ANOVA 'Sepal-1'-'Petal-w' = Variety;
SUBC> MANOVA;
SUBC> Eigen;
SUBC> NoUnivariate.
```

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■ **Results:**

Eigenvalue	32.4704	0.2883	0.0000	0.0000
Proportion	0.9912	0.0088	0.0000	0.0000
Cumulative	0.9912	1.0000	1.0000	1.0000

Eigenvector	1	2	3	4
sepal-l	0.0671	-0.0005	0.2173	0.1481
sepal-w	0.1284	-0.1779	-0.2171	0.0154
petal-l	-0.1814	0.0758	-0.2461	0.0329
petal-w	-0.2331	-0.2331	0.3170	-0.2033

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■ **Results:**

Eigenvalue	32.4704	0.2883	0.0000	0.0000
Proportion	0.9912	0.0088	0.0000	0.0000
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petal-l	-0.1814	0.0758	-0.2461	0.0329
petal-w	-0.2331	-0.2331	0.3170	-0.2033

Zero eigenvalues mean eigenvectors arbitrary, non-interpretable, irrelevant

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■ **Results:**

Eigenvalue	32.4704	0.2883	0.0000	0.0000
Proportion	0.9912	0.0088	0.0000	0.0000
Cumulative	0.9912	1.0000	1.0000	1.0000

Eigenvector	1	2	3	4
sepal-l	0.0671	-0.0005	0.2173	0.1481
sepal-w	0.1284	-0.1779	-0.2171	0.0154
petal-l	-0.1814	0.0758	-0.2461	0.0329
petal-w	-0.2331	-0.2331	0.3170	-0.2033

Zero eigenvalues mean eigenvectors arbitrary, non-interpretable, irrelevant

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■ **Interpretation:**

Eigenvalue	32.4704	0.2883
Proportion	0.9912	0.0088
Cumulative	0.9912	1.0000

1<sup>st</sup> eigenvalue proportionally 0.99 shows most discrimination in just one component (although 3 groups ⇒ 2 crimcoords)

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■ **Interpretation:**

Eigenvector	1	2
sepal-l	0.0671	-0.0005
sepal-w	0.1284	-0.1779
petal-l	-0.1814	0.0758
petal-w	-0.2331	-0.2331

The first eigenvector is (0.07, 0.13, -0.18, -0.23).  
So score on the first crimcoord is  
 $0.07 \times \text{sepal-l} + 0.13 \times \text{sepal-w} - 0.18 \times \text{petal-l} - 0.23 \times \text{petal-w}$

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■ **Interpretation:**

Eigenvector	1	2
sepal-l	0.0671	-0.0005
sepal-w	0.1284	-0.1779
petal-l	-0.1814	0.0758
petal-w	-0.2331	-0.2331

Positive coeffs for sepals  
Negative coeffs for petals  
⇒ Main distinction is sepal size vs petal size

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■ **Producing plots on crimcoords:**

- ◆ Need to multiply data matrix by matrix of eigenvectors

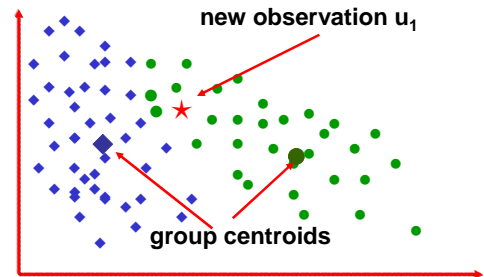
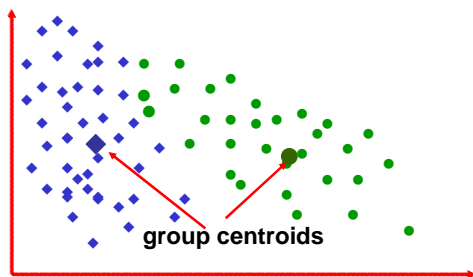
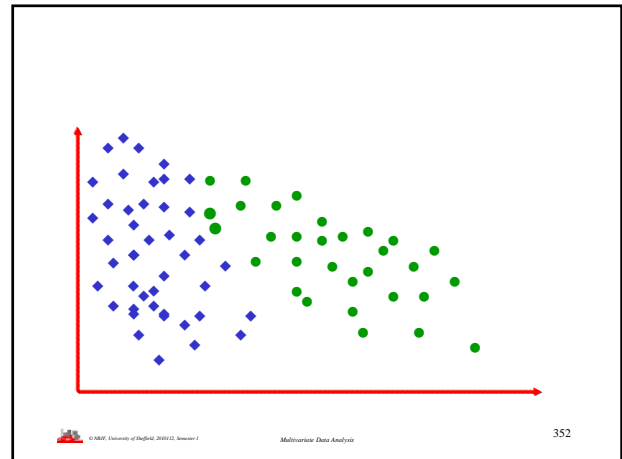
$$\begin{pmatrix} 0.0671 & -0.0005 \\ 0.1284 & -0.1779 \\ -0.1814 & 0.0758 \\ -0.2331 & -0.2331 \end{pmatrix}$$

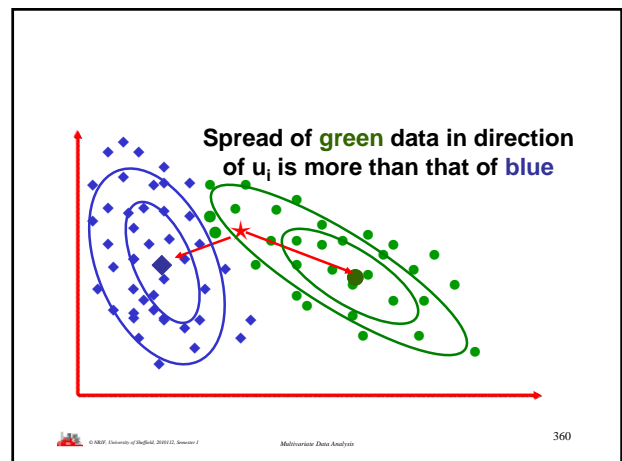
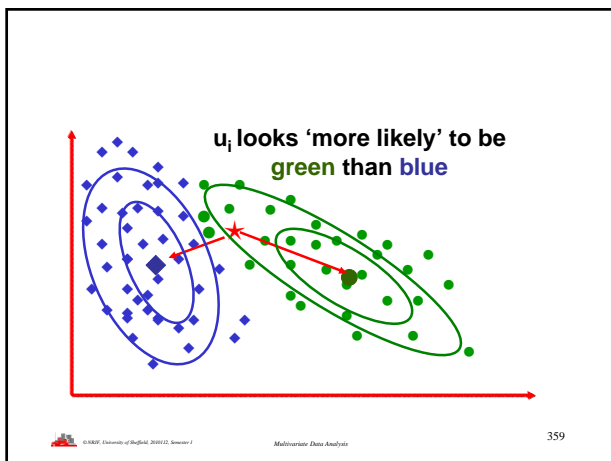
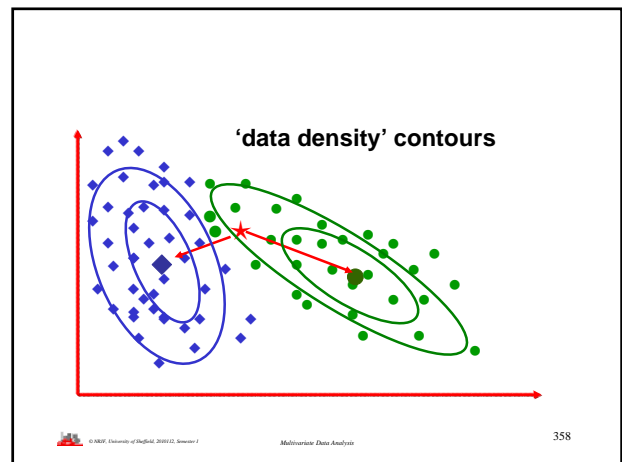
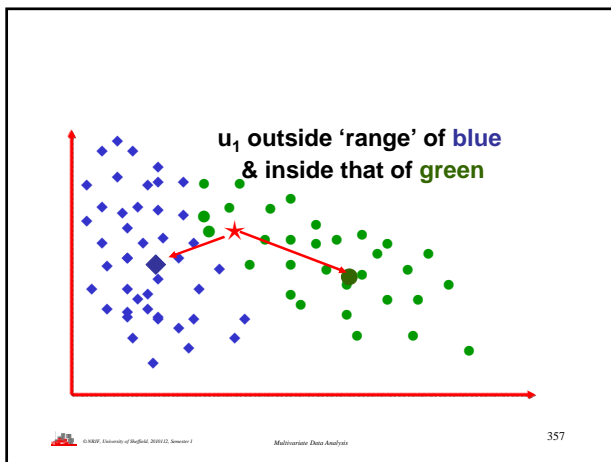
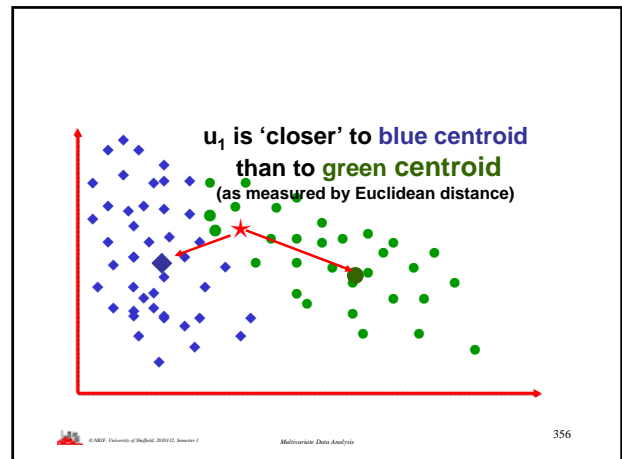
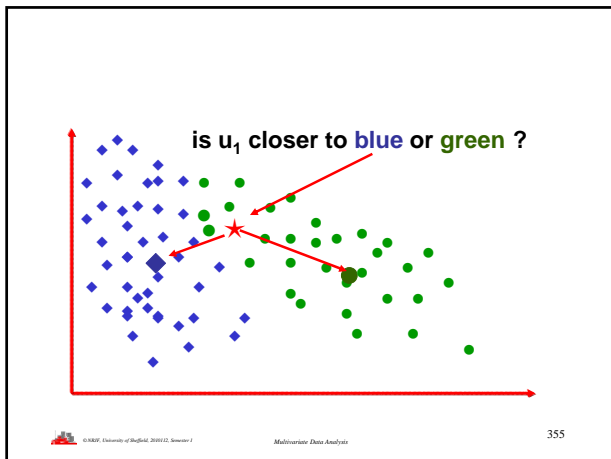
- To multiply data by eigenvectors need to copy columns into matrices

- ◆ Use **Calc > Matrices > Copy** to generate  
 MTB > Copy 'Sepal-l'-'Petal-w' **m1**.  
**Cut and paste eigenvectors to C8 and C9 from session window**  
 MTB > Copy C8 C9 **m2**.  
 MTB > Multiply **m1 m2 m3**.  
 MTB > Copy **m3 c50 c51**.  
 Then **C50 and C51** have the data transformed to 1<sup>st</sup> two crimcoords

■ **Application to Informal Classification**

- ◆ (i.e. Data Analytic Classification)
- ◆ k groups of 'reference' data  $X_i'$ ;  $j = 1, \dots, k$
- ◆ set of r unknown points  $u_j$ ;  $j = 1, \dots, r$
- ◆ allocate each  $u_j$  to the 'closest' group





- Problem arises since **extent of natural variation is different in different directions**

- Sample space is not *isotropic*
- So, need to take differences in direction of variation into account when assessing distance of new observations from group centroids

- Could standardize the data in an analogous way to dividing by standard deviation in univariate data

- Equivalent to recording distance by
 
$$D^2(i) = (u - \bar{x}_i)' S_i^{-1} (u - \bar{x}_i)$$
 if we standardize with respect to each group separately

- or, taking an average of the  $S_i$ , by

$$D^2(i) = (u - \bar{x}_i)' W^{-1} (u - \bar{x}_i)$$

- These measures of distance are termed **generalized distance** or **Mahalanobis distance**

- Historical note:

**Mahalanobis**:- Indian Statistician, key figure in Indian Independence movement, with Gandhi & Nehru founder of Indian Statistical Institute  
main work from ~1935 – 1960

- Instead of

$$D^2(i) = (u - \bar{x}_i)' W^{-1} (u - \bar{x}_i)$$

better to use

$$D^2(i) = (u - \bar{x}_i)' A_t A_t' (u - \bar{x}_i) = [A_t' (u - \bar{x}_i)]' [A_t' (u - \bar{x}_i)]$$

where  $A_t$  is matrix of first  $t$  crimcoords  
(i.e. project data into discriminant space & use Euclidean distance).

Not very different since  $A_t$  crimcoords are eigenvectors of  $W^{-1}B$

### Implementation in R

To plot supplementary data in discriminant space (i.e. on crimcoords)

```
> library(MASS)
> attach(irisnf)
> samp<- c(sample(1:50, 25), sample(51:100, 25),
sample(101:150, 25))
> irisnftraining <- irisnf[samp]
> irisnftest<- irisnf[-samp]
> detach(irisnf)
> attach(irisnftraining)
> iristrain.lda<-
  lda(Variety~Sepal.l+Sepal.w+Petal.l+Petal.w)
> plot(iristrain.lda)
> points(predict(iristrain.lda,
irisnftest[, -5])$x,pch=19)
```

Code for selecting a [random] half of iris data, calculating the crimcoords and then plotting the other half as supplementary points on the crimcoords<sup>65</sup>

### Implementation in R

To plot supplementary data in discriminant space (i.e. on crimcoords)

```
> library(MASS)
• Attach library containing lda(.) function
> attach(irisnf)
• Attach data set
> samp<- c(sample(1:50, 25), sample(51:100,
25), sample(101:150, 25))
• create a vector samp which has 25 random numbers between 1 and 50, 25 in range 51 to 100 & 25 in 101 to 150, i.e. a random half of numbers 1 to 150
> irisnftraining <- irisnf[samp,]
• take a random half of the data, ensuring half of each variety
> irisnftest<- irisnf[-samp,]
• take the other half of the data
```



```

> irisnftraining <- irisnf[samp,]
> irisnftest <- irisnf[-samp,]
• irisnftraining will be used for calculating the lda
  & then want to classify irisnftest
  by plotting these on the crimcoords
> detach(irisnf)
> attach(irisnftraining)
• switch current data set to the test data irisnftraining
> irisnftrain.lda <-
lda(Variety~Sepal.l+Sepal.w+Petal.l+Petal.w)
• calculate the lda using irisnftraining
> plot(irisnftrain.lda)
• plot irisnftraining on the crimcoords
> points(predict(irisnftrain.lda,
+ irisnftest[, -5])$x, pch=19)
• add in the test data irisnftest on the same crimcoords

```

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```

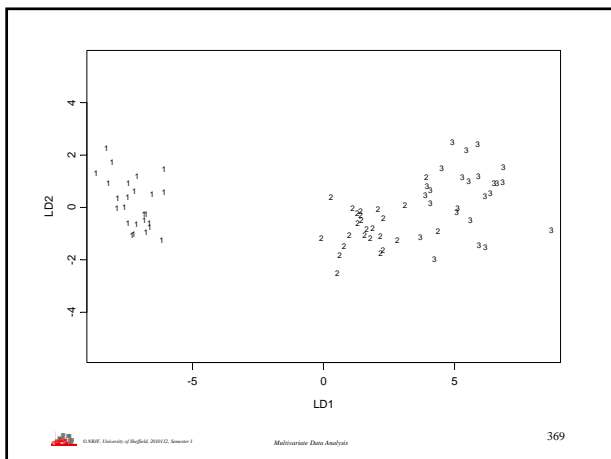
> points(predict(irisnftrain.lda,
+ irisnftest[, -5])$x, pch=19)
WARNING: Point out of bounds: x = -9.346876,
y = 1.295536
• add in the test data irisnftest on the same crimcoords
• note the use of points() to add things to current plot
  & that one of these was outside the plotting area
  • (could have set up the plotting area from
    irisnf with type="n" to get the axes etc)
• note removal of the 5th column from
  irisnftest which has the varieties
• note use of $x to get the coordinates from the predicted data
• note use of pch to get a particular symbol

```

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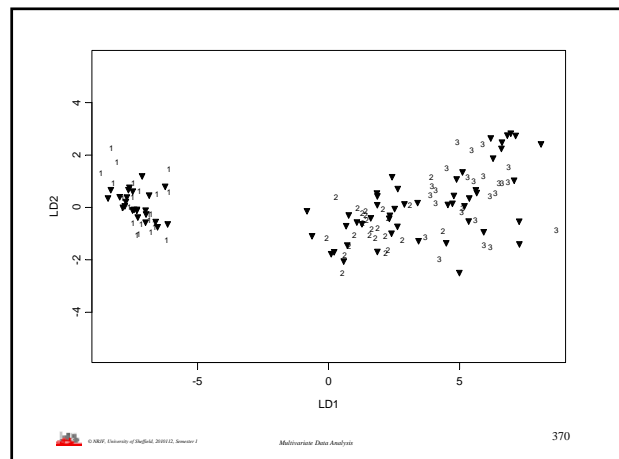
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## Summary and Conclusions

- ◆ Crimcoords highlight differences between **known** groups of observations
  - strictly speaking, displays distorting data slightly (axes not orthogonal but conventionally drawn so)
- ◆ The first Crimcoord is also known as Fisher's Linear Discriminant Function
- ◆ PCA analogy gives interpretations of factor loadings and use of scree plots

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- ◆ Plots on crimcoords gives informal classification of new cases
- ◆ Other classification methods use minimum Mahalanobis distance
- ◆ See Appendix 1 for further illustrations of linear and quadratic discriminant analysis
- ◆ Classification Trees and Neural Networks are related techniques of **supervised learning** (see Appendices 7 & 8) and *Logistic Regression* (see other courses)

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